

IDENTIFYING THE EFFICACY OF CENTRAL BANK INTERVENTIONS: EVIDENCE FROM AUSTRALIA

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Abstract

The endogeneity of exchange rates and intervention has long plagued studies of the effectiveness of central banks' actions in foreign exchange markets. Researchers have either excluded contemporaneous intervention so that their explanators are predetermined, or obtained a small, and typically incorrectly signed, coefficient on contemporaneous intervention. Failing to account for the endogeneity, when central banks lean against the wind and trade strategically, will likely result in a large downward bias to the coefficient on contemporaneous intervention – explaining the negative coefficient frequently obtained.

We use an alternative identification assumption – a change in the intervention policy of the Reserve Bank of Australia – that allows us to estimate, using simulated Generalised Method of Moments (GMM), a model that includes the contemporaneous impact of intervention. There are three main results. Our point estimates suggest that central bank intervention has an economically significant contemporaneous effect. A US\$100 million purchase of the domestic currency will appreciate the exchange rate by 1.3 to 1.8 per cent. This estimate is remarkably similar to the calibration conducted by Dominguez and Frankel (1993c), who themselves noted their estimate was larger than previous empirical findings. Secondly, the vast majority of the effect of an intervention on the exchange rate is found to occur during the day in which it is conducted, with only a smaller impact on subsequent days. Finally, we confirm findings that Australian central bank intervention policy can be characterised as leaning against the wind.

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heteroskedasticity, identification

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1. Introduction

Since the introduction of floating exchange rates, the use and efficacy of intervention in the foreign exchange market has been a controversial topic. Most central banks have at times engaged in frequent intervention and at other times followed a more *laissez-faire* approach to the exchange rate.¹ No doubt the observed disparate range of policies between central banks, and within individual central banks over time, can in part be attributed to the lack of accord on the effectiveness and consequences of central bank intervention. Two key questions remain unresolved: how effective is foreign exchange intervention and, if it is effective, through which channel does it act?

A critical barrier to answering these questions has been overcoming the endogeneity of changes in the exchange rate and intervention. The central hypothesis is that intervention changes the exchange rate. But at the same time, the decision to intervene is not independent of the movements in the exchange rate. Moreover, even once a central bank has decided to intervene, the quantity of currency it buys or sells will typically depend on the response of the exchange rate to its trades.

The literature has typically dealt with the simultaneous equations problem by assuming that the contemporaneous decision of the central bank is independent of the current innovations to the exchange rate. This is a strong assumption. For example, it assumes that the central bank does not change its selling or buying behaviour by assessing the impact its actions have had on the exchange rate. On the other hand, there is strong evidence in stock markets that big players act

¹ Schwartz (2000) suggests that intervention is a dying practice despite the continued use of active intervention by the European Central Bank and Bank of Japan. However, the 18 central banks that responded to a survey reported in Neely (2001) believe it affects the exchange rate. Traders' survey responses in Cheung and Wong (2000) indicate they also believe intervention has an effect on exchange rates.

strategically when they are unwinding large positions. Therefore, why should we expect the same behaviour is not optimal for a central bank?

In this paper we use an alternative identification method to solve the problem of simultaneous equations. We use daily Reserve Bank of Australia (RBA) interventions data over the period 1986–1993, which contains a dramatic change in intervention policy that we use for identification. We show that the estimates we obtain have the correct sign and are significantly larger than those found with more standard methods. This is exactly the direction we would have expected if endogenous variables are an important source of the bias. Further, the vast majority of the effect of an intervention on the exchange rate is found to occur during the day in which it is conducted with a smaller impact on subsequent days. This explains why small effects are usually found when lag values are used in the typical OLS specifications. The major contribution of this paper is to provide some evidence on the contemporaneous effectiveness of intervention. Although our methodology does not indicate the channels through which intervention operates, it provides an improvement from previous estimates obtained in the literature.

We concentrate our analysis on sterilised interventions, but we do not distinguish between secret and public interventions. While undoubtedly this is an important distinction, most and in particular the largest, interventions are public. Certainly, future research should reconsider this issue. In this paper the focus of our attention is the estimation problem. Indeed, we think that the simultaneous equations problem is the crucial aspect limiting our understanding of the effectiveness of policy.

The identification assumption we use is based on the fact that in 1991 the RBA decided to change its policy regarding foreign exchange rate interventions. We interpret this shift in policy as exogenous, which is an important ingredient in our solution to the problem. Other countries have also changed ‘exogenously’ their policies, but typically central banks endogenously respond to the conditions in the market. Again, future research should endogenise the policy decision and extend the present analysis to deal with a more general framework. However, as is argued in Section 3 it is the case that the Australian central bank decision to change its method of intervention was unrelated to other macro events.

The paper proceeds as follows. There is a brief review, in Section 2, of central bank intervention practices and the associated literature. A description of the data

and discussion of RBA intervention follow this, in Section 3. Section 4 outlines the identification and estimation methodology used in this study. The results are presented in Section 5, followed by conclusions.

2. Review of the Literature on Central Bank Intervention

This section briefly reviews the literature on central bank intervention, focusing on the simultaneous relationship between, and temporal behaviour of, exchange rate returns and intervention. For more general and extensive reviews see Sarno and Taylor (2001), Dominguez and Frankel (1993b) and Edison (1993).

Empirical studies, and statements by central banks, suggest that central banks intervene in foreign exchange markets to slow or correct excessive trends in the exchange rate, i.e. they ‘lean against the wind’, and to calm disorderly markets (for example Lewis (1995a) and Baillie and Osterberg (1997b)).² The survey responses of central banks in Neely (2001) suggest that these factors continue to drive the decision to intervene. A recent study for Australia by Kim and Sheen (2002) has similar conclusions. Importantly, when central banks intervene they trade in blocks throughout the day. As Neely (2001) reports, their subsequent trades are conditional on the response of the exchange rate to their earlier trades.

Two main channels have been suggested through which sterilised intervention can affect the level of the exchange rate: the portfolio balance channel and the signalling channel. Intervention changes the exchange rate through the portfolio balance channel if government bonds are imperfect substitutes, and so the change in the reserve asset holdings of a central bank results in private investors revaluing their portfolios of domestic and foreign assets.³ Since interventions are small relative to the stock of outstanding bonds most authors, including Rogoff (1984), have expressed skepticism that intervention could have a large impact through the portfolio balance channel. Not surprisingly, many studies do not find evidence

² Other reasons occasionally cited by central banks include: to target particular exchange rates, or to support other central banks. We do not explicitly consider the impact of intervention on conditional exchange rate volatility. See Rogers and Siklos (2001) for Australia, or more generally Dominguez (1998) and Bonser-Neal and Tanner (1996).

³ The portfolio balance channel requires not only that bonds are imperfect substitutes, but also the failure of Ricardian equivalence.

of this channel, and those that do, such as Dominguez and Frankel (1993a), Evans and Lyons (2001) and Ghosh (1992), suggest that it is weak.⁴

The central bank may also be able to affect the exchange rate by using intervention to credibly disseminate private information about exchange rate fundamentals. This mechanism is described as the signalling channel. Mussa (1981) first suggested this channel with respect to intended changes in monetary policy. The impact of intervention through the signalling channel has often been found to be substantially stronger than through the portfolio balance channel (for example Dominguez and Frankel (1993b)). If the central bank uses intervention to indicate its intended path for interest rates, then this channel can only be an ongoing transmission mechanism if the central bank does follow up with appropriate changes in monetary policy. In this case intervention operating through the signalling channel is not an independent policy tool. Despite the evidence of a signalling channel, Fatum and Hutchison (1999) are unable to find an explicit link between intervention and future monetary policy, while Lewis (1995b) and Kaminsky and Lewis (1996) suggest it occasionally operates in the wrong direction.

Most central banks do not publicly announce their interventions.⁵ In light of the evidence for a signalling channel many authors have questioned this policy.⁶ However, interventions are not completely secret. Dominguez and Frankel (1993b) find that most, and particularly the largest, interventions are reported. Central banks can choose the method of intervention, ranging from direct trades with commercial banks, to indirect trades through brokers, to control the degree of secrecy of their actions.

If intervention operates through the signalling channel, then the exchange rate should react as soon as traders digest the information contained within the intervention. Goodhart and Hesse (1993), Peiers (1997), Dominguez (2003) and

⁴ Dominguez and Frankel (1993c) and Evans and Lyons (2001) both assess the portfolio balance channel for unsterilised intervention and so provide upperbound estimates for the impact of sterilised intervention.

⁵ One exception is the Swiss National Bank as noted by Fischer and Zurlinden (1999).

⁶ Alternatively, Bhattacharya and Weller (1997) and Vitale (1999) develop models of intervention that rationalise secrecy about interventions.

Chang and Taylor (1998) find that interventions that are intended to be visible are typically reported by news services within 10 minutes to 2 hours, by which time it is often no longer ‘news’ to traders. Any effect through the portfolio balance channel is also likely to be rapid as bond holders quickly respond to the change in the relative supplies in the highly liquid market for government securities. Indeed, Neely (2001) reports that the majority of central banks believe the full effects of intervention are reflected in the exchange rate within a matter of hours.

Despite the evidence of the rapid response of the exchange rate to intervention, studies using daily data have often abstracted from the endogeneity of intervention and exchange rate determination by only including lagged intervention (for example, Baillie and Osterberg (1997a) and Lewis (1995b)). While intervention may still have an effect on the days subsequent to the initial trades, omitting the contemporaneous intervention prevents measurement of the immediate impact and is likely to bias other coefficient estimates. Other studies that include contemporaneous intervention, such as Kaminsky and Lewis (1996) and Kim, Kortian and Sheen (2000), typically obtain an incorrectly signed contemporaneous coefficient, suggesting that purchases of the domestic currency cause it to depreciate. Seemingly what they capture is the policy function coefficient that represents central banks’ tendency to ‘lean against the wind’. The insignificant and incorrectly signed coefficients in many previous studies indicate that an accurate estimation of the impact of intervention on the exchange rate must incorporate the contemporaneous effect, and account for the endogeneity between these variables.

3. Reserve Bank of Australia Intervention Data

The estimation in this study uses daily interventions in the foreign exchange market by the RBA over the period from July 1986 to November 1993. This period is chosen because it contains a single distinct change in intervention policy, in October 1991, that allows us to identify the parameters in our exchange rate and intervention system. Additionally, the characteristics of intervention in this sample facilitate accurate estimation. Over the seven years of data (1 930 daily observations) intervention is frequent (48 per cent of days) and often large (up to A\$1.3 billion). Another advantage of our sample is that other central banks do not intervene in the AUD/USD market and so we need only focus on the actions of one central bank.

The data are the daily net purchases of Australian dollars by the RBA and include all transactions made by the RBA, including those on behalf of the Government. Since the RBA can, and does, meet the Government's demand for foreign currency transactions from its own reserves, it has discretion as to the timing of transactions on behalf of the Government. For this reason the inability to exclude transactions on behalf of the Government does not appear to be problematic. Similarly, Neely (1998) argues that not excluding client transactions for the Fed has little influence. Over the sample used, almost all RBA intervention was conducted in the spot market versus the US dollar (Andrew and Broadbent 1994).⁷ Rankin (1998) reports that the RBA has always sterilised its interventions.⁸

The exchange rate is measured as US dollars per Australian dollar (AUD/USD). The intervention and exchange rate returns data are aligned to cover exactly the same 24-hour period, commencing at 9am Sydney time. This avoids inaccurate results from using misaligned data, which have potentially hampered previous studies on central bank intervention. Aligning the intervention and exchange rate returns data is important since central banks can, and do, intervene outside their business hours, as shown in Dominguez (2003), and stated for the RBA by Rankin (1998).⁹ The 24-hour period over which the data are measured will be referred to as a day. The intervention data for any given day will include transactions conducted in any of the major markets around the world on that calendar day.¹⁰

Rankin (1998) outlines five distinct periods of intervention policy used by the RBA since the float of the Australian dollar in December 1983. Table 1 summarises the pattern of intervention over these five episodes. The two periods used in this study

⁷ Over the past decade the RBA has moved to using a combination of spot market transactions and currency swaps, which together replicate a forward, to intervene and sterilise its intervention (see Rankin (1998)).

⁸ Dominguez and Frankel (1993b) and the references contained therein suggest that the Fed, Bundesbank and Bank of Japan have typically only partially sterilised their interventions.

⁹ Given Australian trading hours do not overlap with either London or New York trading hours, and BIS (1993) shows one-quarter of Australian dollar trading occurs in markets outside of Australian trading hours, this is potentially important.

¹⁰ For a few 24-hour periods that end on an Australian public holiday we do not have an observation in the Sydney 9am exchange rate series. Instead, the Australian dollar exchange rate measured at noon EST (2am or 4am Sydney time, depending on daylight saving) by the New York Fed is used.

Table 1: Summary Statistics of Interventions

Regime	Full	I	II	III	IV	V
	Dec-1983 Jul-2000	Dec-1983 Jun-1986	Jul-1986 Sep-1991	Oct-1991 Nov-1993	Dec-1993 Jun-1995	Jul-1995 Jul-2000
Number of interventions						
Sales	1 283	99	780	15	0	389
Purchases	504	230	143	116	0	15
Total number of days	4 219	640	1 338	554	397	1 290
Probability of intervention						
Unconditional	0.42	0.51	0.69	0.24	0	0.31
Conditional on $INT_t - 1 \neq 0$	0.73	0.62	0.80	0.50		0.75
Conditional on $INT_t - 1 = 0$	0.20	0.40	0.44	0.15		0.12
Average intervention (A\$ million)						
Sales	46	9	56	35		36
Purchases	81	16	100	159		285
Absolute value	56	14	63	144		46
Maximum intervention (A\$ million)						
Sale	661	44	661	150	0	286
Purchase	1 585	90	1 026	1 305	0	1 585
Memo items (A\$ million)						
Money base	20 404	11 421	16 206	19 853	22 608	28 995
M1	60 976	22 831	37 261	57 435	74 242	102 921
Notes:	Purchases are of Australian dollars; sample starts 13 December 1983.					
Source:	Authors' calculations					

are the second and third. During the period immediately following the float of the Australian dollar – December 1983 to June 1986 – the RBA allowed the dollar to float freely with only very small interventions. Our study commences using data from July 1986 when the RBA took on a distinctly proactive policy to foreign exchange market intervention. Interventions became more frequent (the RBA was active on 69 per cent of days) and substantially larger (the average absolute intervention was A\$63 million). The change in policy in our sample occurs in October 1991 when the RBA all but ceased to make very small interventions. Large interventions continued to be used as in the preceding era as seen in

Figure 1. As Rankin (1998) puts it, the RBA ‘sought to maximise the impact of its intervention through careful management of their size and timing’. This strategy of not making small interventions drastically reduced the frequency of interventions (to 24 per cent of days), but substantially increased the average size of interventions (to A\$144 million). We do not consider the last two episodes when the RBA did not intervene from December 1993 to June 1995, or when interventions resumed in July 1995.

Figure 1: Proportion of Days with Intervention

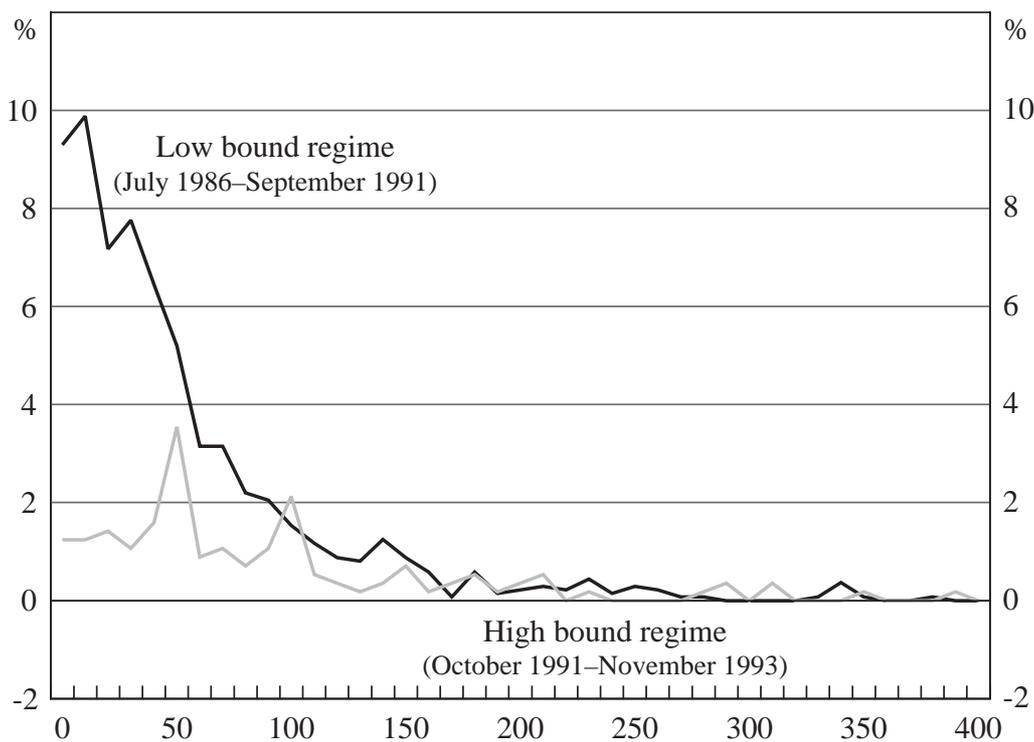


Table 2, drawn largely from Neely (2001), compares the frequency and size of RBA interventions with those of the Fed, the Bundesbank and the Swiss National Bank. Over the period covered in the table, from the early 1980s to the end of the 1990s, the RBA intervened on 42 per cent of days, compared to between 4 and 12.5 per cent for the other central banks. Typical central bank interventions are tiny relative to the huge daily turnover in foreign exchange markets. The daily turnover in the AUD/USD market was estimated to be US\$17.9 billion in 1992 by BIS (1993), although only US\$4.8 billion of this involved a non-financial counterparty.¹¹ While the Australian economy is substantially smaller than the

¹¹ The average (absolute) intervention for the regime including 1992 was just 0.6 per cent of the total average daily turnover.

US and German economies, and the Australian dollar less heavily traded than all three other currencies in Table 2, the RBA interventions are of a comparable size to those of the larger central banks.¹² The large magnitude of some RBA interventions is also confirmed by comparing the maximum interventions to the money base and M1 in Table 1. In the two episodes considered in this study, the largest intervention is seen to be over 5 per cent of the money base, or 2 per cent of M1. Note that while the turnover in the foreign exchange market has increased substantially, and certainly did so through our sample, neither of the hypothesised transmission channels are affected by the volume of turnover.

Table 2: Comparison of Interventions by Central Banks

		Beginning	End	US dollar purchases (US\$ million)		Proportion of days with intervention (Per cent)
				Minimum	Maximum	
US	DEM/USD	01/07/1983	31/12/1998	-797	950	5.6
	JPY/USD	01/07/1983	30/12/1998	-951	722	4.9
Germany	DEM/USD	01/07/1983	31/12/1998	-833	800	12.5
Switzerland	CHF/USD	03/01/1986	29/12/1999	-545	150	3.9
Australia	AUD/USD	13/12/1983	30/07/1999	-932	436	42.3

4. Identification of the Effectiveness of Intervention

In this section we use a simple model of central bank intervention to demonstrate how intervention and the exchange rate interact and the technique we use to solve the identification problem.

A useful starting point to consider the effects of intervention on the exchange rate is a generalised version of uncovered interest parity (UIP)

$$e_t = E^m \{e_{t+1} | \Omega_t\} + i_t - i_t^* + \rho_t \quad (1)$$

¹² The BIS survey lists the AUD/USD currency pair as the ninth most traded at that time. The average daily turnover for the USD/DEM, USD/JPY and USD/CHF was US\$192.2 billion, US\$154.8 billion, and US\$48.8 billion, respectively. Only considering trades with non-financial counterparties the AUD/USD was the fifth most traded currency pair with the three pairs mentioned before respectively US\$26.9 billion, US\$32.3 billion, US\$7.0 billion.

where $E^m \{ \cdot \}$ is the markets' expectation operator, e_t is the exchange rate (value of domestic currency), ρ_t represents possible predictable deviations from UIP (risk premia), and Ω_t is the time t information set, which includes contemporaneous and past interventions, $\{INT_j\}_{j=-\infty}^t \subset \Omega_t$. Substituting forward this relationship, and suppressing other elements of the information set, the exchange rate is

$$e_t = E^m \left\{ \sum_{j=0}^{T-1} [i_{t+j} - i_{t+j}^* + \rho_{t+j}] \middle| INT_t \right\} + E^m \{ e_{t+T} | INT_t \} \quad (2)$$

where e_{t+T} is the exchange rate at some distant time $t+T$. A sterilised intervention to support the domestic currency will increase investors' holdings of foreign bonds and decrease their holdings of domestic bonds. If domestic and foreign bonds are imperfect substitutes investors will require a risk premium ($\rho > 0$) to increase their relative holdings of foreign bonds. Through this portfolio balance channel, for some long-run value for the exchange rate, e_{t+T} , it can be seen in Equation (2) that intervention appreciates the exchange rate.

The signalling channel describes the use of intervention by the central bank to disseminate inside information about exchange rate fundamentals, either interest rates or the long-run value of the exchange rate. Intervention is supposedly a credible communication mechanism because it rebalances the central bank's portfolio in such a way that it profits from resulting changes in the exchange rate and its fundamentals. The central bank will profit from future increases in domestic interest rates if its portfolio has been rebalanced through intervention to support the currency. Intervention can then credibly communicate the bank's intended path for monetary policy, increasing $E \{ i_{t+j} | INT_t \}$. Alternatively, the central bank may use intervention to communicate its view of the 'long-run' value of the exchange rate, so affecting the markets' expectations $E^m \{ e_{t+T} \}$. Since the exchange rate will respond to current bond supplies and current news, both the portfolio balance and signalling channels indicate the exchange rate is a function of contemporaneous intervention, $e_t = e_t (INT_t)$.

As noted earlier, central banks typically state that they intervene to slow or correct excessive trends in the exchange rate and to calm disorderly markets. Indeed, the Governor of the RBA states that the RBA has used intervention in 'circumstances where market imperfections are resulting in overshooting' and that 'intervention can play a useful role in limiting extreme movements in the exchange rate'

(Macfarlane 1998). Since the RBA sterilises interventions, and is acknowledged to allow the Australian dollar to float quite freely, it seems reasonable to assume that intervention is not used as a separate policy tool with independent goals. Rather, intervention is focused on exchange rate outcomes. A simple representation of this policy, as used in Almekinders and Eijffinger (1996), is that the central bank's preferred level of intervention, or shadow intervention, INT_t^* , would minimise squared deviations of the exchange rate from a moving target.

$$L = (e_t (INT_t^*) - \bar{e}_t)^2 \quad (3)$$

Given the central bank is allowing the exchange rate to float, but doesn't want it to move 'too quickly', the target is taken to be a moving average of past values of the exchange rate, $\bar{e}_t = \frac{1}{n} \sum_{j=1}^n e_{t-j}$. The optimal level of intervention will then be given by

$$e_t (INT_t^*) - \bar{e}_t = 0 \quad (4)$$

$$\Delta e_t (INT_t^*) = - \sum_{j=1}^{n-1} \left(\frac{n-j}{n} \right) \Delta e_{t-j} \quad (5)$$

However, central banks do not intervene on every day, and very small interventions are extremely rare. Presumably there are some costs to intervention, possibly because the strength of signals is reduced if they are used too frequently. As a result the central bank only intervenes if the loss function would exceed some benchmark, or equivalently if the shadow intervention exceeds a given threshold, otherwise remaining absent from the market. Actual intervention can then be represented as

$$INT_t = \mathfrak{I} (|INT_t^*| > \overline{INT}) \cdot INT_t^* \quad (6)$$

where $\mathfrak{I}(\cdot)$ is the indicator function. Equations (2), (5) and (6) constitute a system that determines the exchange rate and intervention.

4.1 Set-up of the Estimation System

We generalise the previous framework to include unobservable variables that affect the exchange rate and the central bank's decision for intervention. The reason behind this extension is that we believe that there are factors that are unobservable at daily frequencies that have impact on both variables, such as liquidity shocks, macro shocks, etc. The model used is:

$$\Delta e_t = \alpha INT_t + \gamma z_t + \varepsilon_t \quad (7)$$

$$INT_t = \mathcal{S} (|INT_t^*| > \overline{INT}) \cdot INT_t^* \quad (8)$$

$$INT_t^* = \beta \Delta e_t + z_t + \eta_t \quad (9)$$

where Δe_t is the observed exchange rate return at date t (a positive value is an appreciation), INT_t is the observed intervention (positive values are purchases of the domestic currency), and INT_t^* is a shadow intervention. The estimation procedure includes constants and potentially lags, but this simple version is sufficient to demonstrate the endogeneity problem.

Equation (7) is the reaction of the exchange rate to the central bank intervention. We assume that the exchange rate is affected by two types of shocks: ε_t which is a pure idiosyncratic shock to the exchange rate, which we assume has no direct impact on the intervention decision; and a common shock (z_t) which is assumed to move both the exchange rate and the central bank intervention decision. We explore the interpretation of these shocks below.

Equation (8) is the decision of the central bank to intervene or not. We assume that this decision is made entirely based on the shadow intervention. In other words, if the required intervention is large (larger in absolute terms than some threshold \overline{INT}), then the central bank participates in the market, otherwise it remains absent. Observe that implicitly we assume that if the shadow intervention is larger than the threshold then the central bank intervenes, and its intervention is exactly the shadow one.

Equation (9) determines the shadow intervention. We assume that it is affected by the movements in the exchange rate, by the aggregate or common shock, and by some idiosyncratic shock reflecting innovations to exchange rate policy. If the central bank aims to offset changes in the exchange rate, i.e. lean against the wind,

then β will be negative. Equations (8) and (9) together constitute the central bank's reaction function; where the former reflects the decision to intervene, and the later one determines the quantity or size of intervention.

The policy shock (η_t) is interpreted as innovations in the exchange rate target that are independent of the nominal exchange rate shocks (ε_t) and the common shock (z_t). The idea is to separate idiosyncratic shocks to policy (such as trades on behalf of the government or unwinding of positions) and to the exchange rate (for example, economic fundamentals) from those shocks that we might expect to affect both variables (such as herding, liquidity, or shocks to the exchange rate during periods of high conditional volatility). These common shocks will affect how intervention takes place, and the exchange rate at the same time. We assume that these shocks are i.i.d., with mean zero and variances σ_ε^2 , σ_η^2 , and σ_z^2 . For simplicity in the exposition we have assumed that all the variables have zero mean, but in the empirical implementation it is important to include constants to account for non-zero means.

Finally, the parameter of interest is α . If central bank intervention is effective, then purchases of the domestic currency will appreciate the currency and so α will be positive.

The intuition for this model is that the central bank leans against the wind, so $\beta < 0$, either to slow deviations from trend, or to calm volatile markets. Small changes are tolerable and so the central bank does not bother intervening. On the other hand, if exchange rate returns would otherwise be large, larger interventions would be required to counteract these and will cause the central bank to enter the market. The shadow intervention (INT_t^*) summarises the expected intervention if the central bank were to trade continuously in the foreign exchange rate market.

This simple framework captures the two sources of simultaneity that exist in the data. The first one is the endogenous decision of participation. The second one is the size of the intervention and the change in the exchange rate once the decision of participation has been made. While the first source of bias has been widely acknowledged in the literature, the second has received very little attention. This is understandable. Finding instruments for the first one is hard, but some might be available. For the second one, this is much more difficult.

In this model there does not exist an instrument that can be used to solve the problem of simultaneous equations. More importantly, this bias is likely to be negative, pushing the estimate of α in Equation (7) downward, possibly even negative, explaining most of the results found in the data.

It is important to mention that there are several aspects of central bank intervention that have been oversimplified in this model. First, there is no distinction between public and secret interventions. As was mentioned before, this has received considerable attention in the literature. In this paper we focus on the estimation problem. Second, we do not attempt to distinguish between sterilised and unsterilised interventions as the RBA states that all of its interventions are sterilised.

4.2 Identification through Changes in Intervention Policy

The problem of identification is easily shown by counting the number of unknowns and the number of series we can measure in the model. Under the assumption that we only observe the exchange rate, the size of the intervention, and its timing, then, aside from the means, we can compute only five moments from the data: the probability (or frequency) of intervention; the variance of the exchange rate when there is no intervention; and the covariance matrix when an intervention has taken place. However, in the model there are seven unknown coefficients that explain the behaviour of such variables: the parameters of interest (α , β , and γ), the threshold of intervention (\overline{INT}) and the three variances (σ_ε^2 , σ_η^2 , and σ_z^2).¹³

The standard procedures in the literature use the following assumptions. First, that there are good instruments for the participation decision. Second, that either $\beta = 0$ or $\alpha = 0$ (exclusion restrictions). And third, that the instrument is correlated with η_t but not with z_t . This set of assumptions seems rather strong. Central banks no doubt intervene based on their most recent information set, which includes the change in the exchange rate during the day. Further, the fact that central banks know they have market power, which is the whole rationale for intervening in the first place, collides with the assumption that $\beta = 0$. Central banks should be, and indeed are, strategic in their interventions. The alternative identification

¹³ The estimation of means adds the same number of equations and unknowns to the system, thus, the problem of under-identification remains the same.

assumption that $\alpha = 0$ is similarly problematic in that it implies that central bank interventions don't have any effect on the day during which they are conducted. This contradicts significant circumstantial evidence.

The main contribution of this paper is to relax these set of assumptions and use an alternative identification method that can deal with some of the econometric issues at hand. Obviously, we depend on another set of assumptions. We think those are weaker, in the sense that most of them are already imposed in the standard literature. But this is certainly the first pass at the problem using these alternative methods and further research should extend the present procedure. We discuss the caveats in detail at the end.

Our identification procedure is quite simple; in September 1991, the RBA changed its foreign exchange rate intervention policy. Following the change, the RBA all but ceased to conduct small interventions but continued to undertake larger interventions as before the change. In the model this would be summarised by a shift in \overline{INT} . Effectively, this means that there are two regimes. Under the assumption that the parameters and the variance of the shocks remain the same across both regimes, we have only eight unknowns (one more than before because we have two thresholds) but at least ten moments in the data.

Specifically, the basic model we estimate is the following:

$$\begin{aligned} \Delta e_t &= c_e + \alpha INT_t + \gamma z_t + \varepsilon_t \\ INT_t^* &= c_{INT} + \beta \Delta e_t + z_t + \eta_t \\ INT_t &= \begin{cases} \mathfrak{I} (|INT_t^* - c_{INT}| > \overline{INT}_l) \cdot INT_t^*, & \text{for } t < \hat{t} \\ \mathfrak{I} (|INT_t^* - c_{INT}| > \overline{INT}_h) \cdot INT_t^*, & \text{for } t \geq \hat{t} \end{cases} \end{aligned} \quad (10)$$

Note that in this setup we allow for constants in the mean equations and below we extend the model to also include lags.¹⁴

¹⁴ The lag structure takes into account that the exchange rate can only depend on observable variables. Thus, the lag of intervention used is INT_{t-1} and not INT_{t-1}^* . However, for the shadow intervention equation we allow it to depend on the lag shadow realisation.

We estimate this model with, and without, lags and present both results. When the model is estimated without lags, there are 10 parameters of interest:

$$\{c_e, c_{INT}, \alpha, \beta, \gamma, \overline{INT}_l, \overline{INT}_h, \sigma_\varepsilon, \sigma_\eta, \sigma_z\}.$$

To estimate the model with lags it is necessary to use moments that account for the behaviour of the exchange rate and intervention across days. For consistency, we use the same set of moments to estimate both the models with and without lags. We compute the following moments in each of the regimes: the proportion of days with intervention; the variance of the exchange rate on days with no intervention; and, the variance-covariance matrix on days with intervention. Furthermore, we compute the mean exchange rate return; the mean intervention; the moments related to the serial correlation of the exchange rate; and the probability of consecutive interventions.¹⁵ In total there are 24 moments, greater than the number of parameters, leaving our system over-identified.

We use simulated GMM to estimate the model. The general idea of this procedure can be easily understood by analysing how other techniques estimate the coefficients. For example, when we use Maximum Likelihood (ML) the goal is to estimate the parameters using the mean and the variances. GMM extends that procedure and uses other moments. Simulated GMM is a further generalisation, in which we choose different moments and characteristics from the data, and ‘create’ our own data using our auxiliary model to match those ‘moments’. Indeed, all three techniques use auxiliary models for their estimation. On the one hand, in ML we use multinomial distributions, described only by means and variances. On the other hand, simulated GMM creates its own data within a well-specified model to produce the statistics that we are interested in matching from the population.

In summary, the procedure is as follows: First, we create random draws of 20 000 observations for three uncorrelated shocks with unitary variance. The same set of shock variables are retained for the entire estimation procedure. Second, we simulate the model given some initial conditions and calculate the moments of these simulated data. Third, we compute the ‘distance’ – the sum of the absolute differences – between the population and simulated moments. Finally,

¹⁵ We have run the simulation also including other moments such as the probability of two positive interventions, and the correlation in the quantities of two consecutive interventions. The results were not sensitive to these changes.

we iterate this procedure to search for the coefficients that minimise this distance. Because the probabilities of intervention are two orders of magnitude larger than the variances and covariances we multiply both the intervention and exchange rate data by $\sqrt{10}$ so that the probabilities and variances and covariances are of equivalent magnitudes. This transformation is used to simplify the maximisation procedure.

To calculate the standard errors of the estimates we use the asymptotic distribution of the sample moments. Using the data, we bootstrap the exchange rate, intervention, and probabilities of interventions to produce a sequence of moments (100 of them). Then we estimate the coefficients for each draw of the moments, computing the distribution of our coefficients. Because it is likely that the data are serially correlated, the bootstrap takes this into account.

5. Results

Before we present our results, it is illustrating to demonstrate the problems in estimation that arise from using more traditional methods. As noted earlier, researchers frequently exclude the contemporaneous impact of intervention in an attempt to overcome the endogeneity. To account for the properties of exchange rate data, a GARCH error structure is typically used. Table 3 presents the results from an EGARCH(1,1) model of log changes in the AUD/USD exchange rate over the full sample, and in each of the high- and low-bound regimes. To highlight the impact of the simultaneity bias, we also present regressions that include the contemporaneous intervention. The coefficients on lagged intervention are uniformly small (around 0.005), and frequently the wrong sign, suggesting that RBA purchases of Australian dollars depreciated the exchange rate.

The coefficient on contemporaneous intervention is always significant and incorrectly signed. The estimates are around -0.016 which in absolute value is much larger than the coefficients on lagged intervention. The wrong sign is a direct consequence of the endogeneity. The OLS estimate is a combination of α and β , and likely γ , and can be negative even for positive α if the RBA leans against the wind ($\beta < 0$), or common shocks affect the two equations with different signs ($\gamma < 0$).¹⁶

¹⁶ If there are no lags in the system, the OLS estimate will be $\frac{(\alpha+\gamma)(\alpha\beta+1)\sigma_z^2+\beta\sigma_\varepsilon^2+\alpha\sigma_\eta^2}{(\alpha\beta+1)^2\sigma_z^2+\beta^2\sigma_\varepsilon^2+\sigma_\eta^2}$.

Table 3: EGARCH(1,1) Model of Interventions

	Full sample (07/1986–11/1993)		Low-bound regime (07/1986–09/1991)		High-bound regime (10/1991–11/1993)	
<i>Constant</i>	0.000 (0.000)	0.000 (0.000)	0.000 (0.000)	0.000 (0.000)	0.000 (0.000)	0.000 (0.000)
INT_t			-0.016 (0.001)	-0.020 (0.001)		-0.012 (0.001)
INT_{t-1}	0.004 (0.001)	-0.003 (0.001)	0.009 (0.002)	0.002 (0.002)	0.000 (0.001)	-0.002 (0.001)
INT_{t-2}	0.005 (0.001)	0.002 (0.002)	0.004 (0.002)	0.002 (0.002)	0.005 (0.002)	0.004 (0.002)
<i>Constant</i>	-0.904 (0.115)	-0.807 (0.105)	-1.331 (0.164)	-1.316 (0.154)	-20.122 (0.405)	-19.373 (0.667)
$ \varepsilon_{t-1}/\sigma_{t-1} $	0.286 (0.025)	0.250 (0.024)	0.349 (0.035)	0.333 (0.033)	0.166 (0.049)	0.213 (0.071)
$\varepsilon_{t-1}/\sigma_{t-1}$	-0.053 (0.014)	-0.049 (0.012)	-0.083 (0.019)	-0.085 (0.018)	-0.053 (0.021)	-0.061 (0.031)
$\log(\sigma_{t-1}^2)$	0.934 (0.010)	0.940 (0.009)	0.896 (0.015)	0.895 (0.014)	-0.865 (0.038)	-0.803 (0.064)

Model: $\Delta e_t = c + \phi_0 INT_t + \phi_1 INT_{t-1} + \phi_2 INT_{t-2} + \varepsilon_t$

$$\log(\sigma_t^2) = \omega + \gamma \log(\sigma_{t-1}^2) + \delta \left| \frac{\varepsilon_{t-1}}{\sigma_{t-1}} \right| + \kappa \frac{\varepsilon_{t-1}}{\sigma_{t-1}}$$

5.1 Contemporaneous Effects of Central Bank Intervention

We now move on to the results from our basic model (Equation (10)) that excludes lags. This implies that the central bank target is the previous day's exchange rate, although the rich error structure will capture deviations from this. Later, we extend the model to include lags which imply a target of a longer moving average.

In Table 4 we report the coefficients of interest: α , β , γ , the two thresholds, the constants and the variances. The first column is the point estimate using the simulated GMM methodology and the moments from the data. The second column is the mean of the bootstrapped distribution of the parameter. As mentioned before, we generate 100 sets of moments bootstrapping the residuals of the original data in each of the regimes. The standard deviation is of the bootstrapped distribution of parameter estimates. The fourth column is the quasi t -statistic calculated as

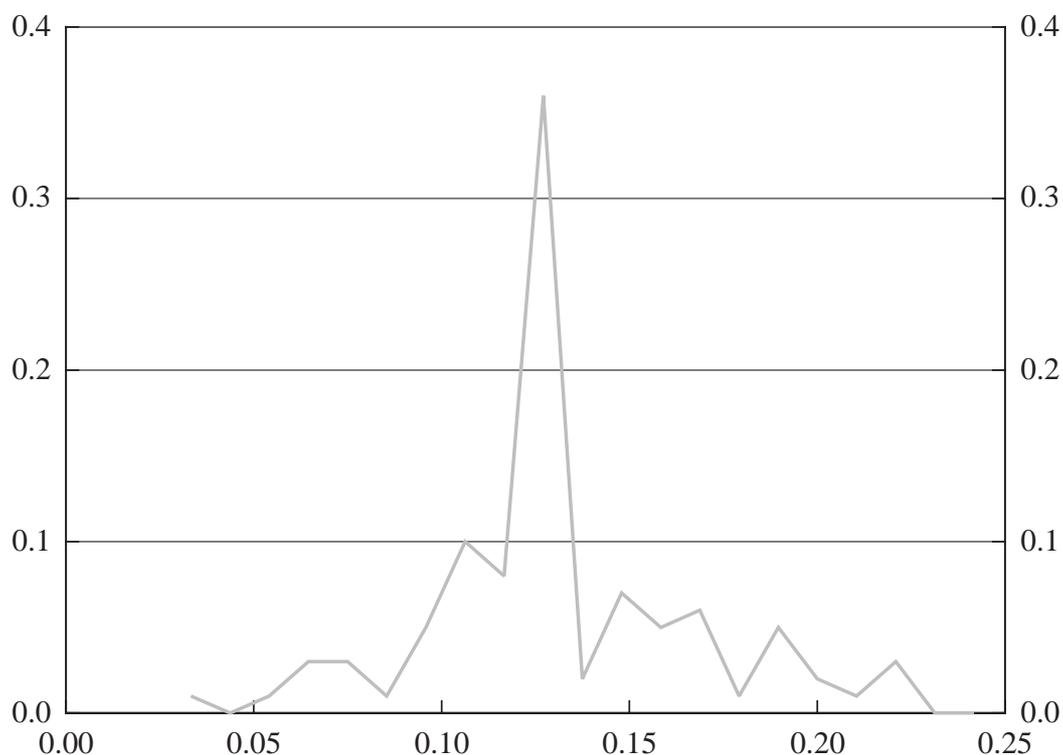
the mean divided by the standard deviation. It is important to highlight that the distributions are not normal, so this statistic should be considered informative but not conclusive in terms of the significance of the coefficients. The fifth and sixth column are the maximum and minimum values of the bootstrapped estimates. The last column is the percentage of the observations that are below zero. This is the statistic that we use to determine the significance of the coefficients. We prefer to look at the mass above and below zero (these are p-values) rather than the quasi t -statistic, given that most of the estimates are not normally distributed.

Table 4: Estimates of the Standard Model

	Point estimate	Mean	Standard deviation	Mean/standard deviation	Maximum	Minimum	Percentage of observations below zero
Exchange rate equation							
α	0.1350	0.1364	0.0372	3.66	0.2313	0.0332	0.0
γ	-0.3042	-0.3054	0.0561	-5.44	-0.1457	-0.4481	100.0
c_e	0.0043	0.0078	0.0188	0.41	0.0904	-0.0859	9.9
Reaction function							
β	-0.0279	-0.0894	0.0952	-0.94	-0.0047	-0.5774	100.0
c_{INT}	-0.0909	-0.1104	0.0552	-2.00	0.0232	-0.3580	99.0
Other parameters							
\overline{INT}_l	0.3160	0.3511	0.0869	4.04	0.9906	0.2112	0.0
\overline{INT}_h	1.8529	1.9767	1.9874	0.99	21.6319	1.2606	0.0
σ_ε^2	0.0345	0.0159	0.0142	1.12	0.0383	0.0000	0.0
σ_η^2	0.4395	0.5025	0.1837	2.73	1.3248	0.2381	0.0
σ_z^2	0.4094	0.4562	0.1825	2.50	1.2857	0.0652	0.0

Our main focus is on the estimate of α . The point estimate is 0.135 and the mean of the distribution is 0.136. The standard deviation is 0.037. Indeed, as can be seen from all 100 realisations, the minimum is 0.033, which implies that there is no mass below zero. Both the quasi t -statistic and the lack of realisations below zero indicate that the estimate is highly significant. The distribution of α is shown in Figure 2.

This estimate of α indicates that intervention has a large effect on the exchange rate. The coefficient on contemporaneous intervention, α (0.136), implies that

Figure 2: Bootstrapped Distribution of α from Basic Model

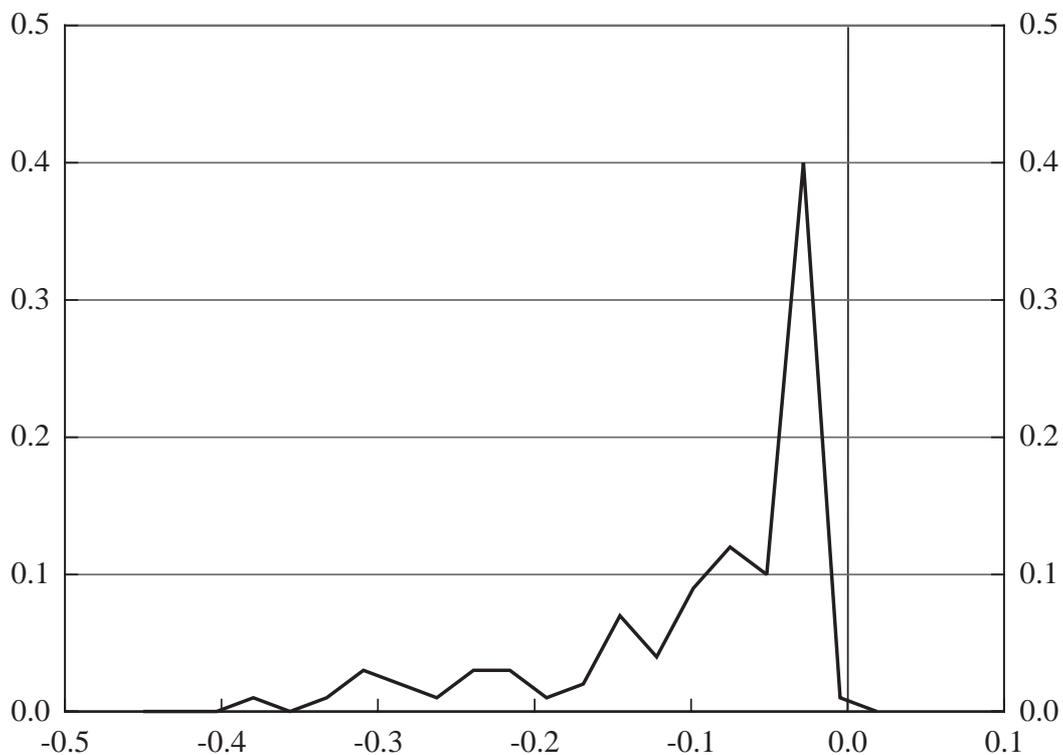
A\$100 million of purchases of Australian dollars is associated (on average) with a 1.36 per cent appreciation of the Australian dollar. If we take the average exchange rate over our sample period to be 0.75, then US\$100 million of purchases of Australian dollars would have appreciated the Australian dollar by approximately 1.81 per cent. This response is larger than most results in the literature but closely related to calibrations obtained by Dominguez and Frankel (1993c). Dominguez and Frankel calculated that US\$100 million of purchases of US dollars would appreciate the US dollar by just under 1.6 per cent.

Our estimate differs from that in Dominguez and Frankel (1993c) in that it is calculated directly from exchange rate and intervention data. In constructing their estimate Dominguez and Frankel need to make several assumptions, such as mean-variance preferences of investors, and use expectations survey data and assets supplies data that are likely to contain measurement errors. The dependence on survey data also requires that they consider exchange rate returns over a longer horizon and so can't estimate the short-run impact of intervention. A significant improvement of our estimation methodology is that we are able to include both contemporaneous intervention and exchange rate returns to address the

simultaneity. Nevertheless, our estimate is close to the one found by Dominguez and Frankel (1993c), and we cannot reject the hypothesis that they are the same.

The point estimate of β is -0.028 , with mean of -0.089 and a relatively large standard deviation of 0.095 . As can be seen in Table 4, the quasi t -statistics suggest that the estimate is not statistically significant from zero. However, demonstrating the importance of the non-normality of the distribution, all realisations are below zero indicating the estimate is highly significant. Figure 3 depicts the distribution of β . As can be seen, the distribution is not normal and there are some realisations that are relatively close to zero, but the distribution has a long left tail.

Figure 3: Bootstrapped Distribution of β from Basic Model



The negative coefficient conforms with our priors, and RBA statements, that interventions are dictated by leaning against the wind. The point estimate implies that in response to a one per cent unexpected depreciation of the exchange rate the central bank will be inclined to lean against the wind, with shadow intervention of a A\$0.28 million purchase of Australian dollars to slow the depreciation.

The parameter γ is precisely estimated; it is negative and significant. The point estimate is -0.304 , almost identical to the mean of the bootstrapped distribution, -0.305 , while the standard deviation is 0.056 . In this case, 100 per cent of the

bootstrapped estimates are negative, implying that the coefficient is significantly different from zero at high levels of significance. The negative sign of γ demonstrates that shocks common to the exchange rate process and the central bank's reaction function simultaneously weaken (strengthen) the currency and increase (decrease) the central bank's desire to support the currency. Again, this supports our priors of central bank policy. These shocks can be interpreted to be shocks to the exchange rate during periods of high conditional volatility which the central bank shows greater inclination to resist. The negative estimate of γ provides an additional reason why failing to account for the simultaneity of exchange rate and intervention determination is likely to bias downward the impact of intervention on the exchange rate.

Finally, the estimates of the thresholds are highly significant and reasonable given what we have observed. Accounting for the scaling of the data, the thresholds imply that in the early period the central bank participates in the foreign exchange rate market when the shadow intervention is larger than A\$3.2 million, while in the later regime it intervenes if the required intervention is larger than A\$18.5 million.

5.2 Dynamic Impact of Central Bank Intervention

We now extend the model to allow for two lags in the structural equations. Doing so provides a richer representation of the policy function and allows us to examine the temporal impact of interventions on the exchange rate.

The expanded model is:

$$\begin{aligned}
\Delta e_t &= c_e + \alpha INT_t + \gamma z_t + \varepsilon_t \\
&\quad + \lambda_{ee,1} \Delta e_{t-1} + \lambda_{ee,2} \Delta e_{t-2} \\
&\quad + \lambda_{eINT,1} INT_{t-1} + \lambda_{eINT,2} INT_{t-2} \\
INT_t^* &= c_{INT} + \beta \Delta e_t + z_t + \eta_t \\
&\quad + \lambda_{INTe,1} \Delta e_{t-1} + \lambda_{INTe,2} \Delta e_{t-2} \\
&\quad + \lambda_{INTINT,1} INT_{t-1}^* + \lambda_{INTINT,2} INT_{t-2}^* \\
INT_t &= \begin{cases} \mathfrak{S} (|INT_t^* - c_{INT}| > \overline{INT}_l) \cdot INT_t^*, & \text{for } t < \hat{t} \\ \mathfrak{S} (|INT_t^* - c_{INT}| > \overline{INT}_h) \cdot INT_t^*, & \text{for } t \geq \hat{t} \end{cases}
\end{aligned} \tag{11}$$

where the additional coefficients are the λ 's. There are 14 coefficients to be estimated. As noted before, there are 24 moments including the across-time statistics. The results from the estimation are shown in Table 5.

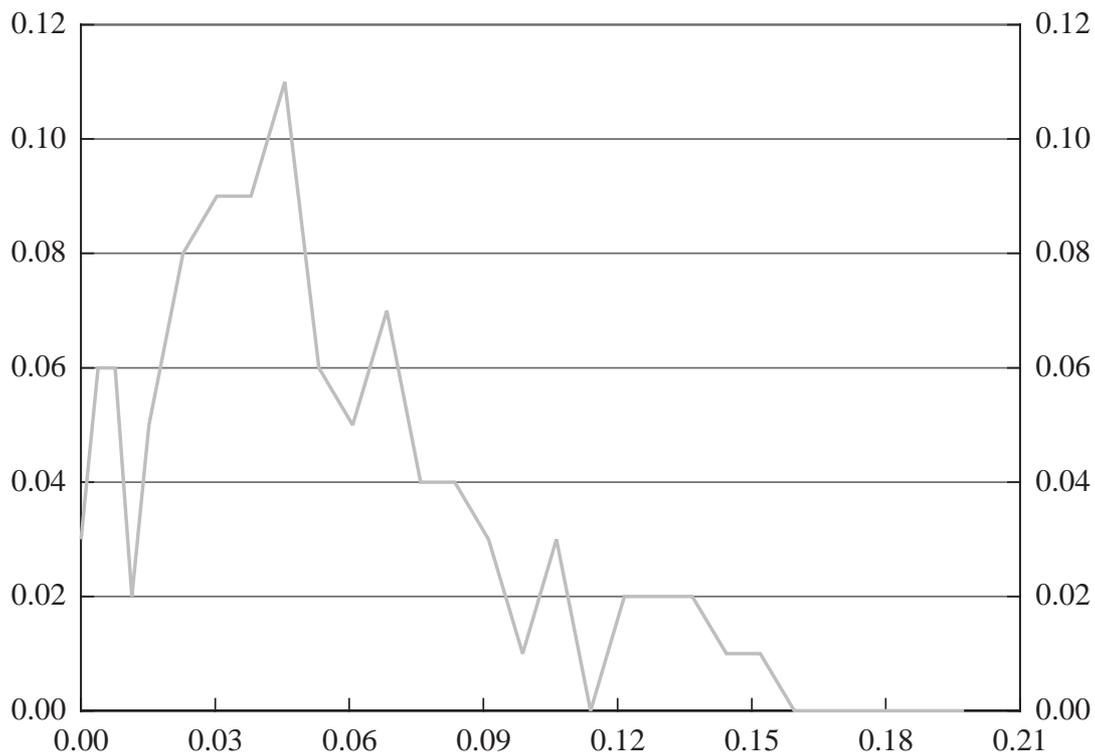
Table 5: Estimates of the Model with Lags

	Point estimate	Mean	Standard deviation	Mean/standard deviation	Maximum	Minimum	Percentage of observations below zero
Exchange rate equation							
α	0.0446	0.0533	0.0375	1.42	0.1579	0.0011	0.0
γ	-0.0612	-0.1400	0.0844	-1.66	0.0021	-0.4130	99.0
λ_{ee1}	-0.0638	-0.0603	0.0348	-1.73	-0.0263	-0.3196	100.0
λ_{ee2}	-0.0379	-0.0385	0.0326	-1.18	0.0085	-0.2962	99.0
λ_{eINT1}	0.0369	-0.0007	0.0351	-0.02	0.0866	-0.0856	65.0
λ_{eINT2}	0.0246	0.0021	0.0321	0.07	0.1060	-0.0844	57.0
c_e	0.0004	0.0100	0.0475	0.21	0.2608	-0.1263	32.0
Reaction function							
β	-0.0657	-0.1538	0.1142	-1.35	0.0000	-0.6292	100.0
λ_{INTe1}	0.0140	-0.0184	0.0186	-0.99	0.0604	-0.0742	87.0
λ_{INTe2}	-0.0030	-0.0189	0.0149	-1.27	0.0760	-0.0558	94.0
$\lambda_{INTINT1}$	-0.2299	-0.0173	0.1499	-0.12	0.2819	-0.6120	43.0
$\lambda_{INTINT2}$	-0.0923	-0.0155	0.1064	-0.15	0.1913	-0.3556	48.0
c_{INT}	-0.0854	-0.1933	0.1150	-1.68	0.0636	-0.6258	98.0
Other parameters							
\overline{INT}_l	0.3253	0.4194	0.1007	4.16	1.0285	0.2500	0.0
\overline{INT}_h	2.6664	1.6935	0.4353	3.89	2.7410	0.7022	0.0
σ_ε^2	0.0128	0.0062	0.0046	1.36	0.0200	0.0000	0.0
σ_η^2	0.3423	0.6430	0.1495	4.30	0.9579	0.2443	0.0
σ_z^2	0.3044	0.4392	0.1283	3.42	0.7659	0.0895	0.0

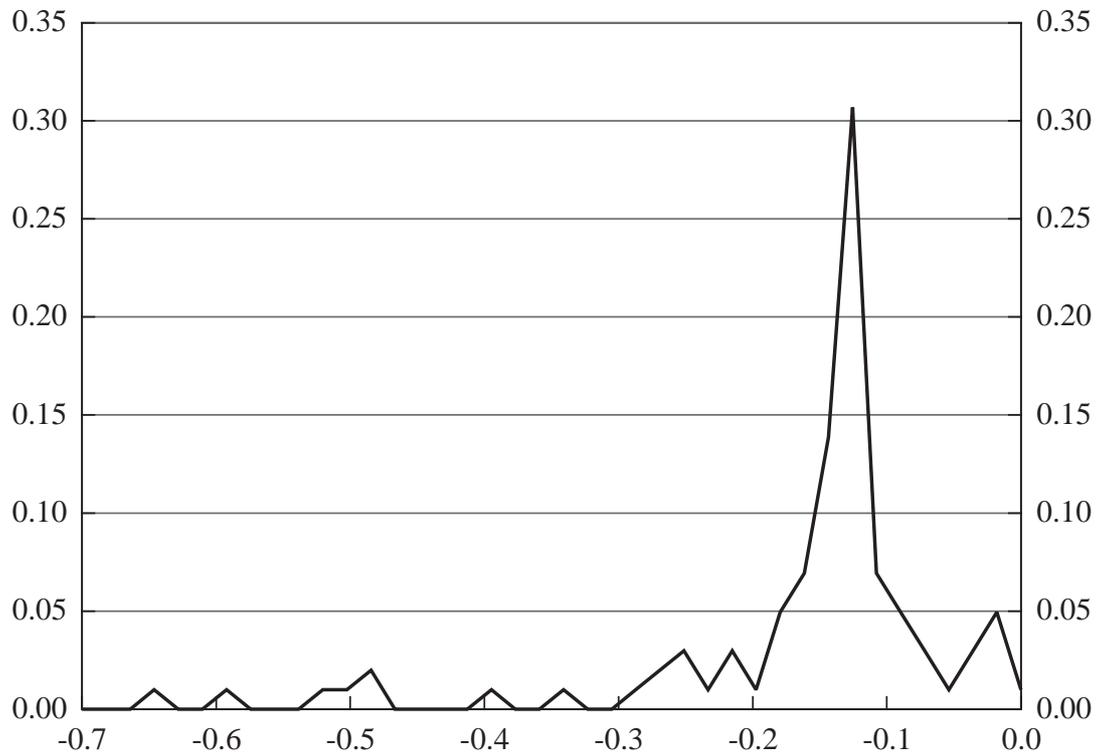
The estimates of α and β have the same signs as those in the base model. The coefficient α is smaller but the coefficients on lagged interventions, $\lambda_{eINT,1}$ and $\lambda_{eINT,2}$, are also positive, indicating that intervention continues to affect the exchange rate on the days subsequent to the intervention. Notably, these subsequent effects are smaller. The largest impact from an intervention occurs on the day it is conducted. The coefficient β is again negative, though larger

in magnitude. The importance of accounting for the endogenous relationship is highlighted by the much smaller role that exchange rate changes from previous days have in determining intervention, as seen by the small coefficients on lagged exchange rate returns. The coefficients on lagged intervention are negative due to the tendency of subsequent interventions to be smaller than the initial intervention. The bootstrapped distributions are depicted in Figures 4 and 5.

Figure 4: Bootstrapped Distribution of α from Model Including Lags



In the model with lags the coefficient on contemporaneous intervention does not measure the total impact of the intervention because there may continue to be changes in the exchange rate for several days. To quantify the impact of intervention it is useful to consider the impulse response function. An additional consideration in the model with lags is that an intervention shock can potentially lead to intervention on subsequent days, both because of the autoregressive and the exchange rate lags in the shadow intervention equation. Whether subsequent interventions occur will depend on the magnitude of the intervention thresholds. We consider two scenarios that bound the range of the exchange rate response. In the first, the intervention threshold is set so that there is no intervention on subsequent days. This is analogous to the traditional presentation of impulse response functions that map a one-time shock and so is comparable to the results

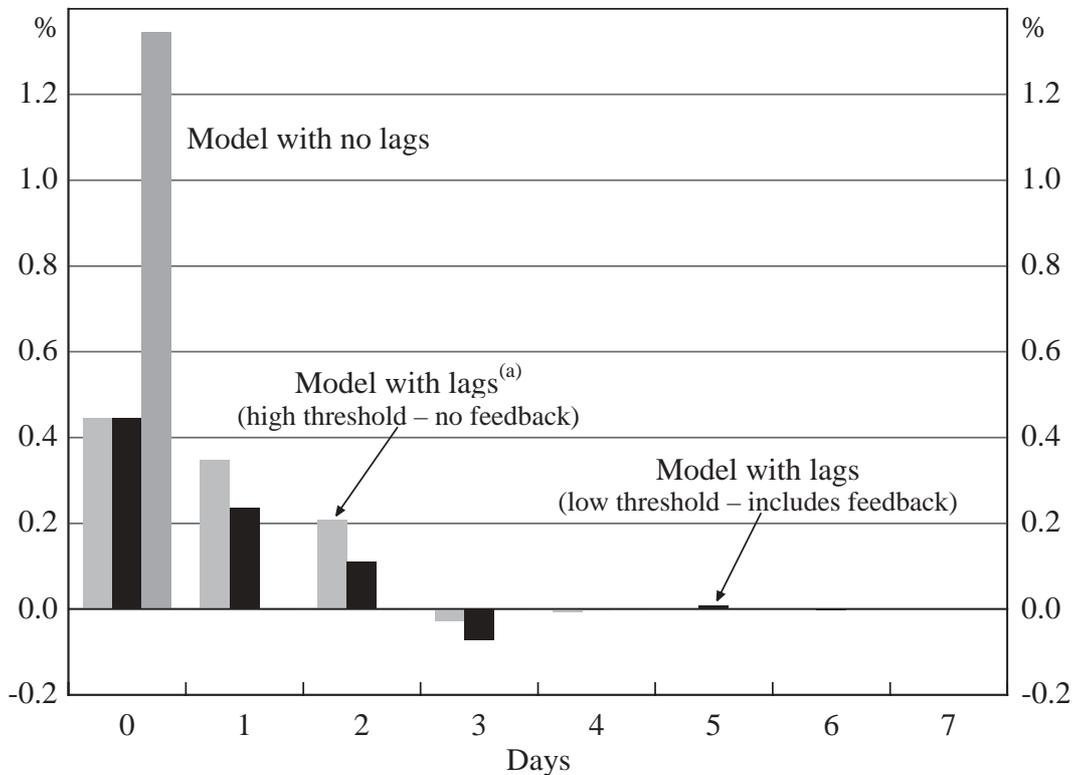
Figure 5: Bootstrapped Distribution of β from Model Including Lags

in the model excluding lags. Secondly, we present impulse responses in which the threshold is effectively zero and so any non-zero shadow intervention is actually conducted.

Figure 6 shows these two impulse response functions along with the one for the base model without lags. Since the coefficients in the policy function on lagged interventions are negative, using a zero threshold for intervention implies that the central bank will partially unwind its intervention on subsequent days. Empirically this does not occur – the central bank almost never intervenes on opposite sides of the market on subsequent days. This highlights the importance of the interaction of the two components of the central bank policy function. The negative coefficients on lagged intervention result in a muted response of the exchange rate when the low threshold is used. Both scenarios demonstrate that the largest impact of intervention comes on the day it is conducted.

Figure 6: Impulse Response

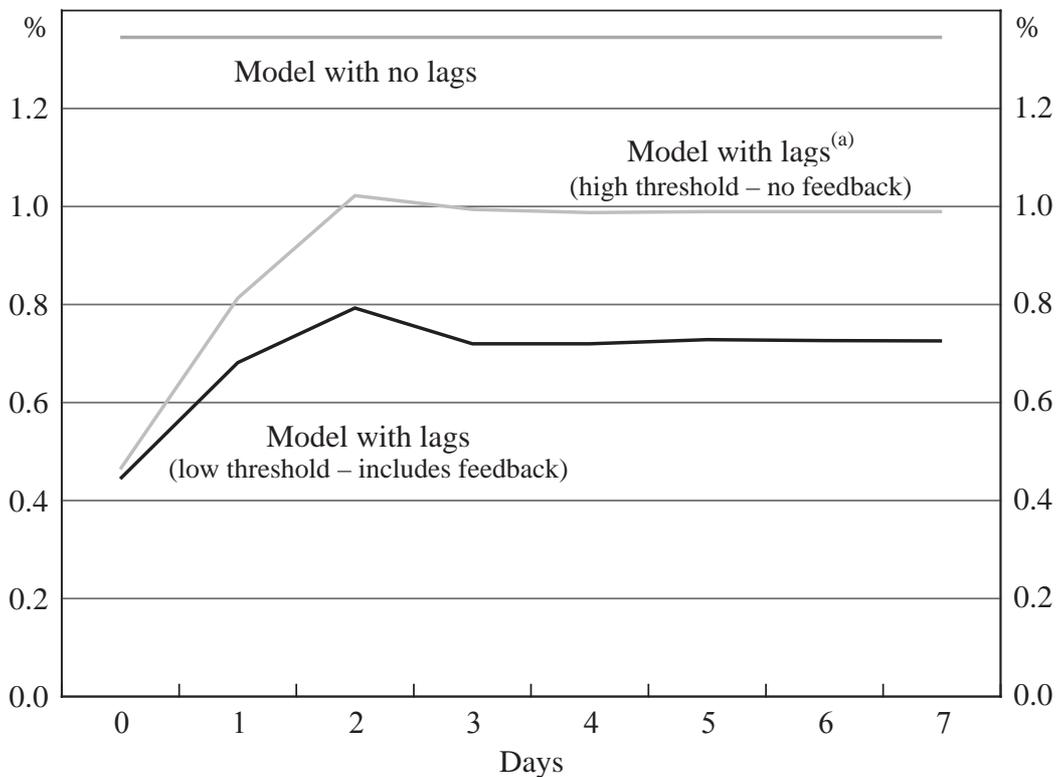
Change in exchange rate from a A\$100 million intervention shock



Note: (a) Restricts intervention equation so there is no intervention following the initial day

The cumulative impact of intervention is shown in Figure 7. Note that the short lag structure used implies that intervention has a permanent effect on the exchange rate. A long lag structure would be needed to capture the unwinding of the effects of intervention. Including many lags is not practical for the estimation method we employ and so we are unable to assess how long the effects of intervention last beyond a few days. The cumulative response of the exchange rate to intervention, after two to three days, is similar for the models including and excluding lags. In the model with lags A\$100 million (US\$100 million) will appreciate the exchange rate by 1 per cent (1.3 per cent), slightly less than the 1.4 per cent (1.8 per cent) predicted by the model excluding lags.

Figure 7: Impulse Response
 Increase in the level of the exchange rate from a
 A\$100 million intervention shock



Note: (a) Restricts intervention equation so there is no intervention following the initial day

5.3 Caveats and Further Research

Before concluding, it is worth spending time discussing several caveats that should be addressed in future research. First, our estimation methodology does not distinguish through which channel intervention affects the exchange rate. Given the relative shifts in bonds in private hands is small, the portfolio balance channel is likely to play an insignificant role. Rather, the signalling channel, and possibly market microstructure effects, would seem to be responsible – but this is a conjecture, rather than a result.

Second, as it stands our methodology requires that the change in policy is truly exogenous. Finding such exogenous changes in policy for other countries may not be straight forward. However, it is possible to model the threshold in the participation equation as a function of macro variables, the exchange rate,

and even the second moments of the endogenous variables. The procedure of estimation would be exactly the same as the one described here if the change in policy is thought of as a shift in the coefficients in that equation. Furthermore, the participation decision could be rationalised as a switching Markov regime where the transition probabilities are functions of the endogenous variables. The estimation of these models is beyond the scope of the present paper and is left for future research. Nevertheless, it is important to mention that in this paper we have emphasised the identification issue, and our specification is a reduced form representation of these models where the shift in policy would be reflected in changes in the intervention decision equation.

Finally, our methodology assumes that the coefficients, other than the threshold, are stable across regimes. This seems to be a reasonable assumption in the current study but may be more difficult to justify in other contexts. This criticism is akin to the application of the Lucas critique to models of monetary policy. Our framework is not entirely exempt from this critique, and therefore, conclusions from this analysis are subject to the caveat of the strength of these assumptions.

6. Conclusion

The endogeneity of exchange rates and intervention has long plagued studies of the effectiveness of central banks' actions in foreign exchange markets. Researchers have either excluded contemporaneous intervention, so that their explanators are predetermined, or obtained a small, and typically incorrectly signed, coefficient on contemporaneous intervention. Failing to account for the endogeneity, when central banks lean against the wind and trade strategically, will likely result in a large downward bias to the coefficient on contemporaneous intervention – explaining the negative coefficient frequently obtained.

This paper uses a novel identification assumption, a change in RBA intervention policy, that allows us to estimate a model that includes the contemporaneous impact of intervention. We use simulated GMM to estimate the model. There are three main results. Our point estimates suggest that central bank intervention has a substantial effect. We find that a (sterilised) purchase of US\$100 million of Australian dollars by the RBA would be associated with an appreciation of between 1.3 and 1.8 per cent. These estimates are remarkably similar to that in Dominguez and Frankel (1993c), even though our estimation methodology

is completely different, depending on only exchange rate and intervention data. Second, an intervention is shown to have its largest effect on the exchange rate on the day in which it is conducted, with smaller effects on subsequent days. This finding has not previously been demonstrated in the literature due to the problem of endogeneity, and confirms the beliefs of central banks of the immediacy of the effects of intervention. Finally, we confirm findings that Australian central bank intervention policy can be characterised as leaning against the wind.

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